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1 Original Research

Q2  
3 Validating a Time Series of Annual Grass Percent Cover in the Sagebrush Ecosystem<sup>☆</sup>

Q3  
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## A B S T R A C T

We mapped yearly (2000–2016) estimates of annual grass percent cover for much of the sagebrush ecosystem of the western United States using remotely sensed, climate, and geophysical data in regression-tree models. Annual grasses senesce and cure by early summer and then become beds of fine fuel that easily ignite and spread fire through rangeland systems. Our annual maps estimate the extent of these fuels and can serve as a tool to assist land managers and scientists in understanding the ecosystem's response to weather variations, disturbances, and management. Validating the time series of annual maps is important for determining the usefulness of the data. To validate these maps, we compare Bureau of Land Management Assessment Inventory and Monitoring (AIM) data to mapped estimates and use a leave-one-out spatial assessment technique that is effective for validating maps that cover broad geographical extents. We hypothesize that the time series of annual maps exhibits high spatiotemporal variability because precipitation is highly variable in arid and semiarid environments where sagebrush is native, and invasive annual grasses respond to precipitation. The remotely sensed data that help drive our regression-tree model effectively measures annual grasses' response to precipitation. The mean absolute error (MAE) rate varied depending on the validation data and technique used for comparison. The AIM plot data and our maps had substantial spatial incongruence, but despite this, the MAE rate for the assessment equaled 12.62%. The leave-one-out accuracy assessment had an MAE of 8.43%. We quantified bias, and bias was more substantial at higher percent cover. These annual maps can help management identify actions that may alleviate the current cycle of invasive grasses because it enables the assessment of the variability of annual grass – percent cover distribution through space and time, as part of dynamic systems rather than static systems. © 2018 The Society for Range Management. Published by Elsevier Inc. All rights reserved.

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## 42 Introduction

43 Annual grass percent cover in the sagebrush ecosystem is highly variable both interannually and spatially. After these grasses complete their lifecycles each year, they become highly flammable and create beds of fine fuel capable of spreading fire (Whisenant, 1990; Balch et al., 2013). Because of the spatial and temporal variability of annual grasses, the annual extent and potential impact of these beds of fuel over large geographic areas are unknown without yearly maps that estimate annual grass percent cover. We developed a satellite-based time series (2000–2016) of maps that estimate annual grass percent cover in the

sagebrush ecosystem of the western United States and identify the extent of these beds of fuel. These maps can serve as a tool that helps land managers and scientists understand the ecosystem's response to weather variations, disturbances, and management in the context of annual grass variability. This understanding can occur because annual grasses are positively correlated with precipitation (Bradley and Mustard, 2005; Pilliod et al., 2017), and the combined effect of fire and grazing leads to reduced resistance to annual grass invasion through altered dynamics of other biotic factors (Condon and Pyke, 2018). As annual grass extents expand in sagebrush ecosystems, the associated biodiversity loss and continuity of fine fuels results in grass–fire cycles (Brooks et al., 2004) that increase the threat to adjacent sagebrush communities, increase the danger to human-built structures, reduce air quality, and compromise grazing and recreational resources. This new fire regime induces the replacement of native plant species with invasive plants (D'Antonio and Vitousek, 1992; D'Antonio, 2000; Brooks et al., 2004), which causes the displacement of wildlife species, reducing their populations (Connelly et al., 2011). In the past 30 yr, wildfires have caused more widespread damage in western ecosystems than occurred

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under the historical fire regime (Connelly et al., 2011; Schoennagel et al., 2017). Balch et al. (2013) determined that fire frequency, size, and duration have increased substantially in areas infested with cheatgrass, a dominant invasive annual grass in the sagebrush ecosystem. With climate change and climate-driven vegetation change playing important roles in the potential transformation of fire regimes (Liu and Wimberly, 2016), the lengthening of the fire season in the western United States is likely. Kerns and Day (2017) discovered that in experimental plots in the Blue Mountains ecoregion within the Malheur National Forest, Oregon, autumn burns led to higher cover of cheatgrass than areas that experience no burns (controls) or spring burns. These findings make knowledge of the extent and magnitude of the annual grass invasion essential, both its interannual and spatial variabilities, especially as endogenous and exogenous influences change. Therefore, while validating the accuracy of a time series of annual grass – percent cover maps with ground-truth data can be problematic, the validation is important for determining data limitations.

In this study, we used 7-day best pixel composites of enhanced Moderate Resolution Imaging Spectroradiometer (eMODIS) normalized difference vegetation index (NDVI) data that were available weekly (the eMODIS NDVI data are described in more detail in the Methods section and can be downloaded at <https://earthexplorer.usgs.gov/>). Consistent and frequent acquisitions of spatially explicit remotely sensed data can mitigate some challenges of using remotely sensed data (Jenkerson et al., 2010; French et al., 2013) and be used to build time series datasets that allow monitoring of many diverse phenomena. eMODIS NDVI senses relatively quick-changing ecological processes (Wylie et al., 2012; Browning et al., 2015) and has done so over long periods at relatively low cost because of its fine-scale temporal acquisition schedule, use of data from a long-flying satellite, and broad coverage. eMODIS NDVI data can also be used to measure the variability of annual vegetation's response to weather, disturbances, and/or management (Wylie et al., 2012). Maynard et al. (2016) discovered that in a time series (2000–2012), MODIS NDVI predicted vegetation biomass better than Landsat 5 largely due to MODIS's high temporal resolution composites. A time series of remotely sensed data is valuable, in part, because it allows the establishment of an average value for each mapped unit. Comparisons between that average and its time series can elicit valuable information when one or more periods deviate substantially from normal (Boyte et al., 2018).

Remotely sensed – based ecological modeling projects generally rely on field data or its derivatives, and field data are generally difficult and expensive to obtain and a challenge to directly associate with remotely sensed – based results (Bradley et al., 2018). Validating a time series of ecologically and remotely sensed – based results with field-based datasets can be problematic for at least two reasons. First, independent field data are scarce and may not exist in a study area for all, or even some, years in a time series (Browning et al., 2015). Second, the spatial resolution of the remotely sensed data and the spatial representation of field plots may be incongruent and require either the spatial manipulation of one dataset to match the other or the acceptance of spatial resolution differences between datasets. Either of these circumstances influence validation efforts when comparing satellite data with field data. When conducting remotely sensed ecological studies in sagebrush ecosystems, the problems of spatial incongruence between the remotely sensed data and the field plots can be exacerbated (Maynard et al., 2016) because, in their native states, these ecosystems can have highly heterogeneous vegetation patterns with areas of substantial bare ground that produce mixed satellite reflectance signals. Therefore, the larger the spatial footprint of the individual pixels in the remotely sensed data, the more likely the differences will be substantial between the remotely sensed data and the field data (Browning et al., 2017).

The goals of this study are twofold: 1) contribute to the understanding of the historical annual grass invasion in the sagebrush ecosystem and 2) visually demonstrate the spatial and interannual variability of

annual grass percent cover in this study's time series. In the process of reaching these two goals, we accomplish three objectives: 1) develop a time series (2000–2016) of yearly maps that describe the relative abundance and extent of annual grass percent cover in the sagebrush ecosystem; 2) describe and illustrate the annual grass – percent cover mapping model and its results; and 3) report the relative accuracy of this time series. Within the context of these objectives, we hypothesize that, in already invaded areas, the interannual variability in the time series of maps will closely follow seasonal precipitation patterns and the overall accuracy of the maps will show a mean absolute error (MAE) rate of < 10%. We establish these hypotheses on the basis of our understanding of previous work that describes cheatgrass response to highly variable precipitation (Bradley and Mustard, 2005; Pilliod et al., 2017) and drivers of environmental resistance to exotic brome grasses (i.e., cheatgrass [*Bromus tectorum* L.] and red brome [*Bromus rubens* L.]), such as elevation and soil moisture and temperature regimes (Chambers et al., 2016).

## Methods

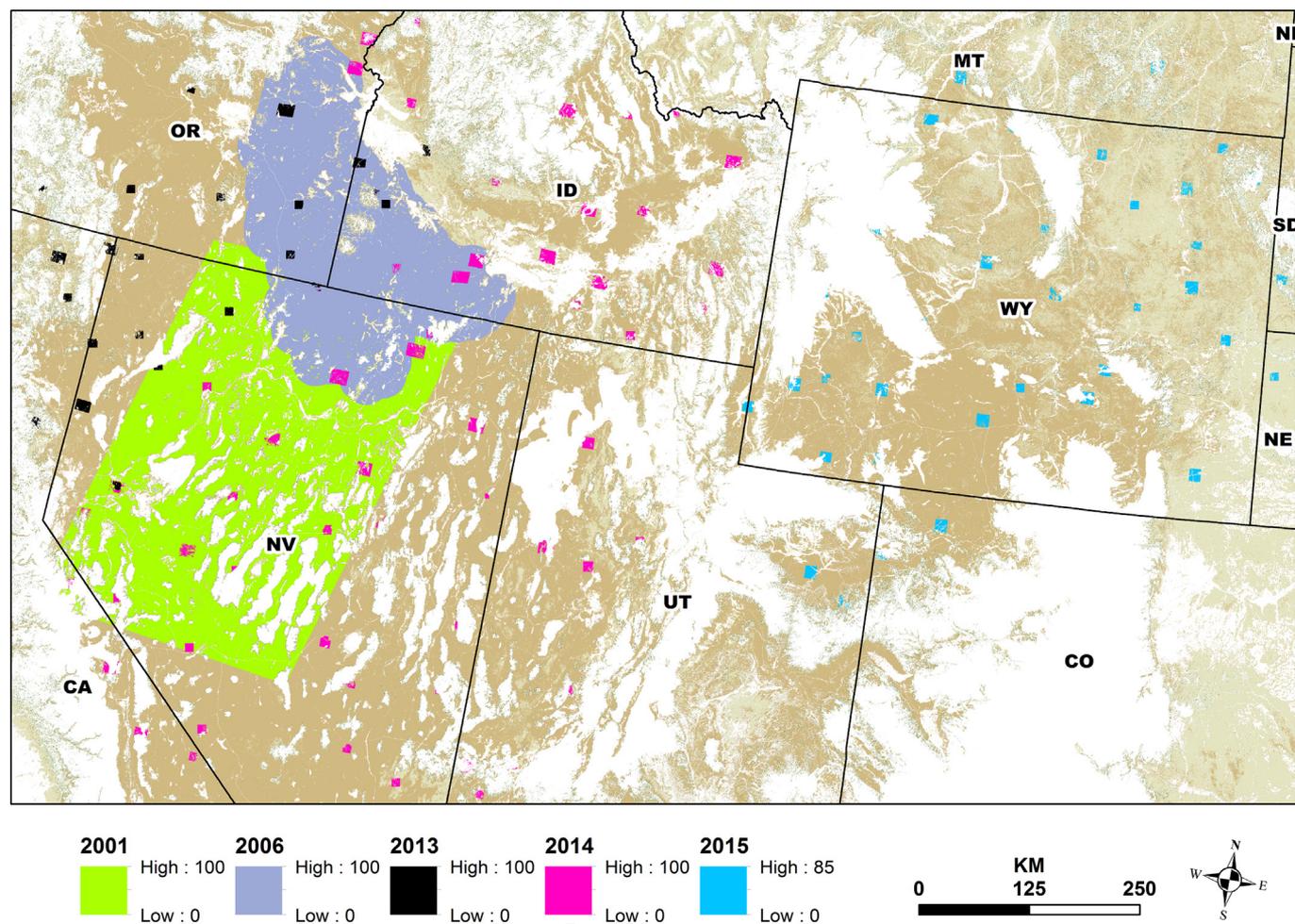
### Study Area

The study area encompassed about 1.3 million km<sup>2</sup> of the western United States, including all of Wyoming and parts of 10 other states (Fig. 1). All or part of 20 ecoregions fell within the study area's boundary, including all of the Northern Basin and Range, the Central Basin and Range, and the Snake River Plain (Commission for Environmental Cooperation, 2009). This study focused on land dominated by shrub and grassland/herbaceous vegetation (National Land Cover Database ([NLCD], <http://www.mrlc.gov/nlcd2001.php>) at or below 2 250-m elevation, and this included approximately 52.5% of the mapped area. A mask covered the other 47.5% of the area. In previous work, we mapped average cheatgrass percent cover at < 2% at an elevational range of 1 750–2 000 m (Boyte et al., 2015b, 2016). Consequently, we set a 2 250-m threshold to focus our study on likely areas of annual grass invasion while allowing for expansion of the future annual grass envelope. The study area's 30-yr (1981–2010) average precipitation equaled 416 mm with a range from 46 mm in the lower, drier areas to 2 890 mm on higher peaks. The 30-yr temperature averages ranged from –0.38° C to 14° C (PRISM Climate Group, <http://prism.oregonstate.edu>). Elevations ranged from –72 m to 4 357 m with a mean of 1 818 m (North American Vertical Datum of 1988). Much of the area was a shrub steppe environment that historically was dominated by sagebrush (*Artemisia* spp.). The sagebrush coexisted with perennial grasses like bluebunch wheatgrass (*Pseudoroegneria spicata* [Pursh] A. Love), Idaho fescue (*Festuca idahoensis* Elmer), and Sandberg bluegrass (*Poa secunda* J. Presl) and annual invasive grasses like cheatgrass (*Bromus tectorum* L.), ventenata (*Ventenata dubia* [Leers] Coss.), annual red brome (*Bromus reubens* L.), and medusahead (*Taeniatherum caput-medusae* [L.] Nevski) (West and Young, 2000). Other common woody species included rabbitbrush (*Chrysothamnus* Nutt.), winterfat (*Krascheninnikovia* Guldenstaedt), greasewood (*Sarcobatus* Nees), shadscale (*Atriplex confertifolia* [Torr. & Frem.] S. Watson), and fourwing saltbush (*Atriplex canescens* [Pursh] Nutt.) (Wiken et al., 2011). Other herbaceous species included Thurber's needlegrass (*Achnatherum thurberianum* [Piper] Barkworth), squirreltail (*Elymus elymoides* [Raf.] Swezey), western wheatgrass (*Pascopyrum smithii* [Rydb.] A. Love), and needle and thread (*Hesperostipa* [Elias] Barkworth) (Wiken et al., 2011).

### Data

#### Dependent Variables

We accessed three spatially explicit datasets with 30-m spatial resolution as reference data. This included a north-central Nevada cheatgrass percent cover dataset (~2001) and an Owyhee Upland annual



**Figure 1.** The study area showing the general spatial distribution of training data locations. Each location is colored based on the yr the training data represent. Multiple pixels reside within each location, but every potential pixel is not selected as a training point. The values associated with each yr show that yr's training data percent cover range. We harvested training data only from locations classified by the National Land Cover Database (NLCD) as shrub or grassland/herbaceous and at or below 2 250 m elevation. A mask (white) covers all other areas. The size and shape of each training data location varies on the basis of the pixels within that location that meet the NLCD and elevation parameters. The 2011 NLCD serves as a backdrop.

198 grass index (2006) (Nevada Natural Heritage Program, <http://heritage.nv.gov/>). We also used percent cover estimates of annual herbaceous  
 199 vegetation (2013–2015) datasets. These datasets used WorldView  
 200 data and were developed by a team of researchers, field technicians,  
 201 and Global Information System/remote sensing experts associated  
 202 with the US Geological Survey (USGS) NLCD project (Xian et al.,  
 203 2015). Because we utilized reference datasets from sources that quan-  
 204 tify cheatgrass, annual grass, and annual herbaceous vegetation we  
 205 refer to the data this study produced as annual grass – percent cover  
 206 maps. We used the cheatgrass and annual grass datasets in two previous  
 207 publications (Boyte et al., 2015a, 2016) where we described the datasets  
 208 and their accuracy. In brief, these two datasets were developed using  
 209 2001 Landsat 7 Enhanced Thematic Mapper Plus (ETM+) or 2006  
 210 Landsat 5 Thematic Mapper (TM) satellite data combined with geo-  
 211 physical variables along with data from more than 650 field plots  
 212 (Peterson, 2005, 2007). The 2001 cheatgrass dataset experienced a cor-  
 213 relation coefficient ( $r$ ) of 71% and a root mean squared error (RMSE) of  
 214 9.1% (Peterson, 2005). The 2006 annual grass dataset had an RMSE be-  
 215 tween 10% and 16%, with 75% of field plots within 14% of the field mea-  
 216 surements (Peterson, 2007). The annual herbaceous datasets were  
 217 developed using a multi-scaling approach that integrated sample data  
 218 from field plots with high-resolution (~2 m) WorldView-2 and  
 219 WorldView-3 satellite data with 8 spectral bands (Xian et al., 2015). Al-  
 220 gorithms that predict annual herbaceous percent cover were developed  
 221 with rule-based regression-tree software, and then the algorithms were  
 222

223 applied to a mapping application that generated 2-m spatially explicit  
 224 estimates. The 2-m estimates were spatially averaged to a 30-m spatial  
 225 resolution using the degrade tool in ERDAS Imagine. This tool averages  
 226 the original pixels that compose the new, larger spatial resolution  
 227 pixels. These 30-m datasets were then, as were the 30-m cheatgrass  
 228 percent cover and annual grass index data, spatially averaged using a  
 229  $7 \times 7$  focal mean and resampled to 250-m to match the spatial resolu-  
 230 tion of the eMODIS NDVI data.

#### Independent Variables

231  
 232 Independent variables were chosen that help quantify annual grass  
 233 percent cover. These variables had to be spatially explicit and cover  
 234 the entirety of the study area. They included satellite, topographic,  
 235 soils, climatic, land cover, and disturbance datasets. The satellite data  
 236 were generated from 17 yr (2000–2016) of 250-m eMODIS NDVI data  
 237 and used to develop four derivative variables for each year: mean grow-  
 238 ing season NDVI (Spring GSNs), mean summer NDVI (Summer GSNs),  
 239 annual grass indices, and estimated start of season spring growth. The  
 240 NDVI product used the MODIS red (620–670 nm) and near-infrared  
 241 (NIR) (841–876 nm) bands in an equation (Eq. (1)) (Jensen, 2005)  
 242 that measures dynamic vegetation greenness. The eMODIS NDVI data  
 243 consisted of 7-d best pixel composites derived from daily data where  
 244 a minimum-value-composite algorithm identified the best pixel in  
 245 each 7-d period by filtering through input surface reflectance of poor  
 246 quality, negative values, low view angles, clouds, and snow cover

(Jenkerson et al., 2010). These composites mitigated problems caused by clouds, shadows, off-nadir fields of view, and atmospheric effects (Jenkerson et al., 2010). After acquisition, the composites were temporally smoothed to further mitigate residual cloud effects. The 7-d temporal resolution of the eMODIS NDVI data captured the phenological dynamics of rapidly developing annual grasses and, when combined with this data source’s ability to mitigate problems inherent to remotely sensed data, made it a strong choice to monitor annual grass percent cover on an annual time step.

$$NDVI = \frac{\rho_{nir} - \rho_{red}}{\rho_{nir} + \rho_{red}}, \text{ where } \begin{cases} \rho_{red} = \text{red reflectance} \\ \rho_{nir} = \text{near infrared reflectance} \end{cases} \quad (1)$$

The annual grass indices were a function of Spring GSNs and Summer GSNs (Eq. (2)) (Kokaly, 2011). Each variable was spatially and temporally dynamic and reflected the variation in weather, disturbances, management, and site characteristics at each pixel. The start of season spring growth was a phenologically driven variable used only to identify each pixel’s dynamic starting point (the eMODIS NDVI weekly composite) of Spring GSN’s integration period (Boyte et al., 2015b). We incorporated an independently developed 250-m remotely sensed phenology dataset derived from eMODIS NDVI, start of season time (SOST) (<http://phenology.cr.usgs.gov/index.php>), into the model as an independent variable. To account for fire disturbances in our model, we included a 250-m time-since-fire dataset that used Monitoring Trends in Burn Severity data (<https://www.mtbs.gov/>). Elevation data and their derivatives, steep slope with aspect and a compound topographic index, were produced from 30-m data available from The National Map (<https://nationalmap.gov/elevation.html>). Slopes exceeding 8.5% were classified as steep. North-facing and south-facing slopes were defined by azimuths from 315 degrees to 45 degrees and 135 degrees to 225 degrees, respectively. Digital soils data included 30-m available water capacity and soil organic matter from the POLARIS website (<http://stream.princeton.edu/POLARIS/>). To define precipitation zones in our study area, we resampled 30-yr averages (1981–2010) of PRISM (<http://prism.oregonstate.edu>) precipitation data at 800 m spatial resolution to 250 m using bilinear interpolation. We used one categorical dataset, the 30-m 2011 National Land Cover Database, to help stratify the model at areas classified as shrub or herbaceous. We spatially averaged all 30-m datasets and then resampled them to 250 m to match the eMODIS NDVI data.

$$\text{Annual Grass Index} = \frac{\text{Spring GSN} - \text{Summer GSN}}{\text{Spring GSN} + \text{Summer GSN}}, \quad (2)$$

where  $\begin{cases} \text{Spring GSN} = \text{integrated growing season NDVI} \\ \text{Summer GSN} = \text{integrated summer NDVI} \end{cases}$

287  
288 *Developing the Rule-Based Regression-Tree Model*

The dependent and independent variables described earlier provided information on training cases for our model. We entered these variables into rule-based, regression-tree software (<https://www.rulequest.com/>) and used them to develop a spatially explicit model of annual grass – percent cover estimates. The regression-tree software stratified rules by relating the dependent variable to the independent variables. Each rule had an associated algorithm that was used to develop an estimation for all pixels that fit that specific rule. We reduced the number of rules and removed some variables from our model (Table 1). The fewer rules employed by the model, the more generalized the model estimates. The more rules employed by the model, the more specific the model estimates, and the more likely the model would be overfit (Gu et al., 2016). More rules could also have added intermittent spatial artifacts to mapping outputs, artifacts that reflect the spatial

**Table 1** Driving variables for the annual grass – percent cover model. Frequency (%) of use is shown for each variable used to establish rule conditions and the associated linear regression models. Dashes indicate that a variable was not used. The model was constructed of 5 committees and 27 rules. t1.1–t1.5

Driving variable	Rule conditions	Linear regression model	
Spring GSN	78	95	t1.6
Elevation	78	93	t1.7
Available water capacity	50	69	t1.8
30-yr precipitation	45	77	t1.9
Start of season time	33	73	t1.10
Annual grass index	16	68	t1.11
Time since fire	15	41	t1.12
National land cover database	13	–	t1.13
Soil organic matter	11	77	t1.14
Summer GSN	9	93	t1.15
Compound topographic index	–	35	t1.16
South facing steep slope	–	14	t1.17
North facing steep slope	–	13	t1.18

Based on 9 randomizations: training data  $R^2 = 0.76$ . Mean absolute error =  $5.79 \pm 0.03$ . Test data  $R^2 = 0.74$ . Mean absolute error =  $5.84 \pm 0.11$ . Ten-fold cross validation  $R^2 = 0.74$ . Several variables were initially applied to the model but omitted because they caused excessive spatial artifacts in the maps. These variables include a latitude proxy, Major Land Resource Area, LANDFIRE environmental site potential, 30-yr temperature maximums, and 30-yr temperature minimums. t1.20–t1.25

pattern of specific independent variables and not what was on the ground. Some modelers remove independent variables that add spatial artifacts to their maps. We tested the number of model rules at multiple values between 100 and 25, and we found that 27 rules optimized the model relatively well and left its mapping outputs mostly free of spatial artifacts. We trained the annual grass model on 33 746 randomly stratified points that were spaced through time (Table 2) and spatially distributed throughout the unmasked portion of the study area (see Fig. 1) where herbaceous/grasslands or shrub lands were the dominant land cover (<http://www.mrlc.gov/nlcd2001.php>). We harvested the training points from the 5 yr of data described in the dependent variables subsection and built a model robust to dynamic conditions like weather, disturbances, and management. Training a synoptic model with points that represented a variety of conditions encountered through time and space reduced model extrapolation and created better model estimates (Jung et al., 2009; Gu et al., 2012). 303–318

*Developing the Time Series of Annual Grass Maps* 319

The model developed by the regression-tree software for the annual grass estimates was applied using a mapping application, MapCubist. USGS Earth Resources Observation and Science (EROS) Center computer scientists developed MapCubist using publicly available source code provided by RuleQuest (<https://www.rulequest.com/>). This code used GDAL (<http://www.gdal.org/>), an open-source raster processing library to produce an application capable of reading a list of rasters, applying the rule-based, linear regression equations to the independent variables and producing output estimates for each year in the time series. All rasters of independent variables represented the entire study area extent. 320–329

**Table 2** The number of training points by yr. t2.1–t2.2

Yr	Points	
2001	4804	t2.3
2006	4418	t2.4
2013	4970	t2.5
2014	10061	t2.6
2015	9493	t2.7

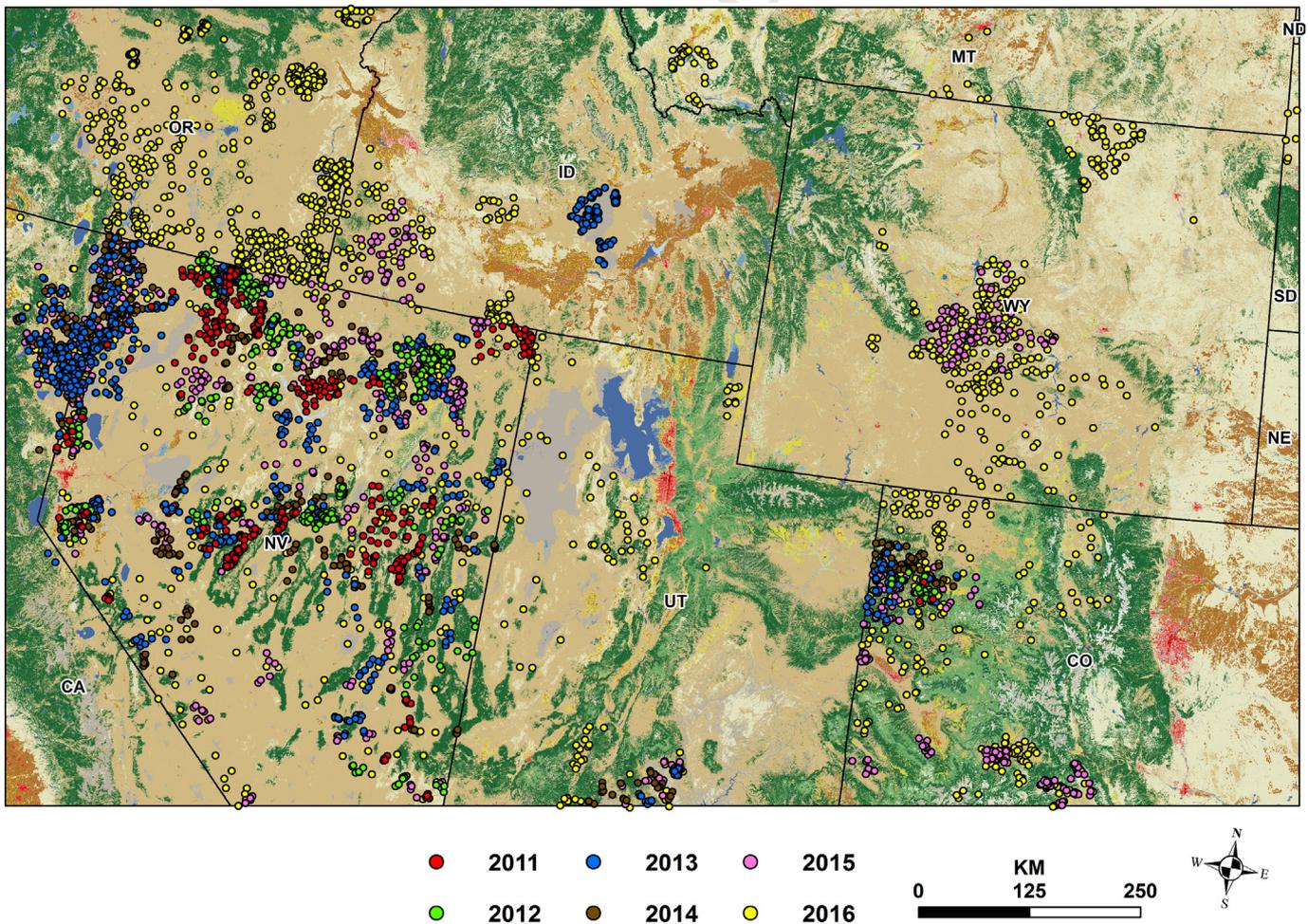
## 331 Model Evaluation

332 The regression-tree software produced model estimates based on  
 333 the associations between the dependent and independent variables  
 334 and generated model accuracy metrics. These metrics included the cor-  
 335 relation coefficient and an MAE rate. Several user-controlled parameters  
 336 changed the model structure, and these parameters included the num-  
 337 ber of committee models used iteratively to improve substantially in-  
 338 correct estimates, percentage of model extrapolation allowed, and  
 339 maximum number of rules the model used to stratify the data. Once  
 340 we optimized the model's parameters (Gu et al., 2016), we set the  
 341 regression-tree software to begin with a random seed and then con-  
 342 ducted nine bootstrap model runs. We averaged the accuracy metrics  
 343 of the nine model runs for overall model accuracy. Separately, we gen-  
 344 erated a model with a 10-fold cross validation where each of the 10  
 345 folds was withheld sequentially as test data. This validation process  
 346 also output a correlation coefficient and an MAE rate. We converted  
 347 the correlation coefficient to an  $R^2$ .

## 348 Assessment Inventory and Monitoring Data

349 We downloaded several yr (2011–2016) of Bureau of Land Manage-  
 350 ment (BLM) Assessment Inventory and Monitoring (AIM) data that co-  
 351 incided with our study period and area ([https://gis.blm.gov/](https://gis.blm.gov/AIMdownload/layerpackages/BLM_AIM_Terrestrial.lpk)  
 352 [AIMdownload/layerpackages/BLM\\_AIM\\_Terrestrial.lpk](https://gis.blm.gov/AIMdownload/layerpackages/BLM_AIM_Terrestrial.lpk)) (Fig. 2). AIM  
 353 data are designed “to quantitatively assess the condition, trends,  
 354 amount, location, and spatial pattern of natural resources on the na-  
 355 tion's public lands” (<https://landscape.blm.gov/geoportal/catalog/AIM/>

AIM.page). Herrick et al. (2017) described the collection of AIM data 356  
 where a line-point intercept method was used with a pin to measure 357  
 vegetation percent cover and composition. Plot transect lengths typi- 358  
 cally extended from 25 to 50 m, and plot designs took several forms in- 359  
 cluding a spoke, parallel lines, or a single straight or curved line. We 360  
 compared the AIM data to our mapping estimates, calculated the corre- 361  
 lation coefficient, MAE rate, and normalized root mean square error 362  
 (nRMSE) for each individual year and for all years combined (Table 3). 363  
 The nRMSE is a dimensionless statistic that measures model fit with 364  
 no regard for a dataset's range and allows comparison between multiple 365  
 RMSE calculations (Homer et al., 2012). Two incongruities exist be- 366  
 tween the AIM data and the mapped estimates of annual grass percent 367  
 cover. First, the AIM process uses an “any hit” technique—any time a pin 368  
 is dropped, every plant the pin touches is recorded as a hit and included 369  
 in the percent cover calculation, irrespective of whether the hit is in the 370  
 upper or lower canopy. This can result in vegetation cover hypotheti- 371  
 cally being recorded at > 100%. Passive radiometers (visible, infrared, 372  
 and shortwave infrared) on satellites respond primarily to the charac- 373  
 teristics of the top canopy layers with low or no sensitivity to the 374  
 lower canopy levels in dense vegetation. Mapped estimates did not ex- 375  
 ceed 100%. Second, and likely more important, spatial resolution differ- 376  
 ences existed between the AIM data and mapped estimates, which is a 377  
 common problem when validating remote sensing data (Bradley et al., 378  
 2018). Spatially, the AIM-transect plots represented approximately 379  
 4.5–15.2% of a 250-m pixel (a 250-m pixel represents 6.25 ha, so 380  
 AIM-transect plots represent  $\approx 0.28–0.95$  ha), depending on transect 381  
 lengths in the plot design. Consequently, in many cases, AIM annual 382



**Figure 2.** Assessment Inventory and Monitoring (AIM) plot locations delineated temporally by color overlaid on the 2011 National Land Cover Database. We compared AIM data with corresponding mapped estimates of annual grass percent cover to assess the maps' accuracies.

**Table 3**  
 Comparing Bureau of Land Management Assessment Inventory and Monitoring (AIM) plots to annual grass – percent cover mapped estimates. The AIM plots represent from 4.5% to 15.2% of a 250-m pixel, depending on the transect length in the plot design. The spatial resolution disparity between datasets likely affects the statistics that describe the relationship. The units are percent based on AIM plot data.

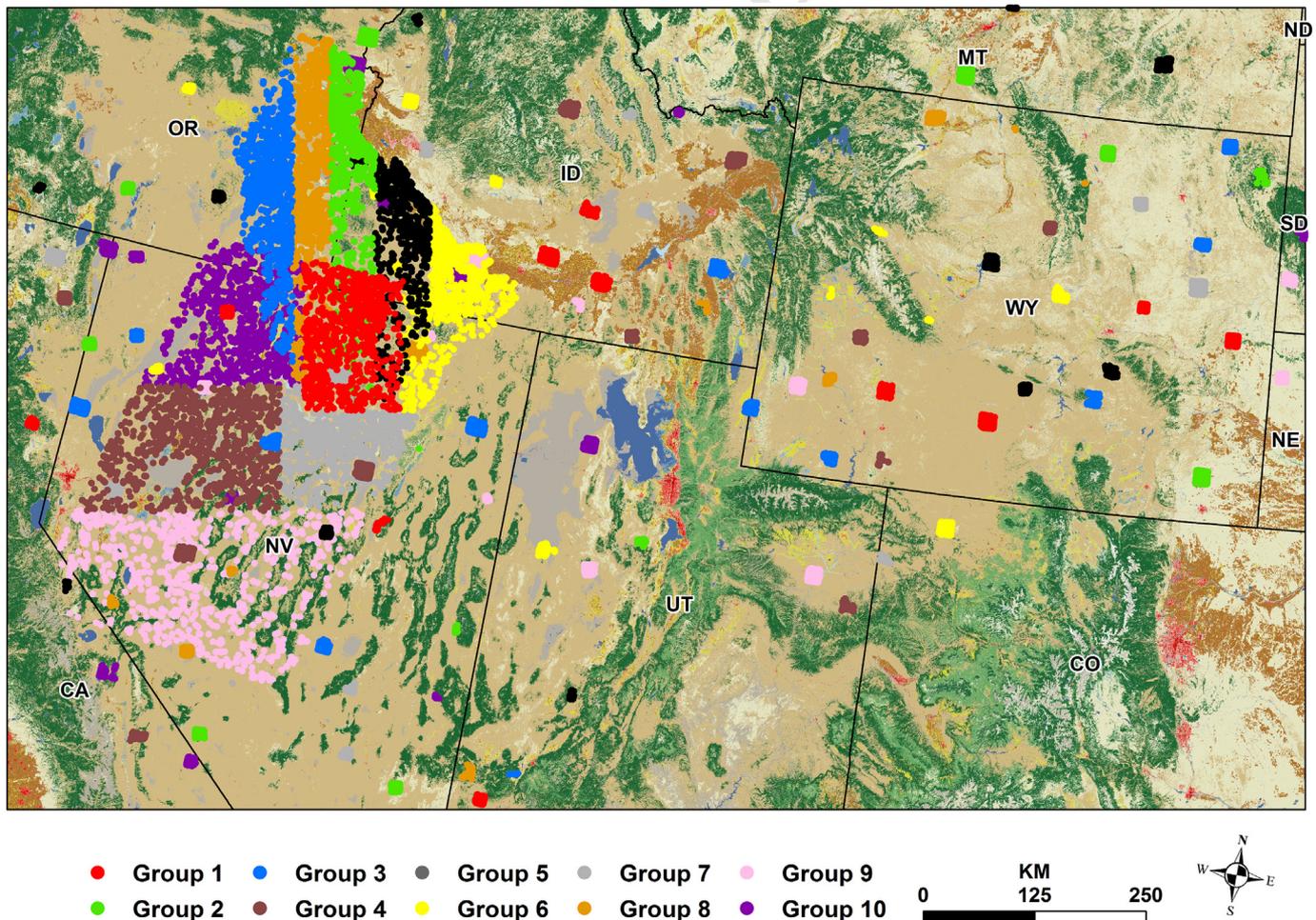
Yr	<i>r</i>	MAE (%)	nRMSE (%)	<i>n</i> =
2011	0.60	15.69	0.23	332
2012	0.55	14.92	0.19	367
2013	0.39	7.89	0.17	712
2014	0.50	12.15	0.17	569
2015	0.49	12.44	0.20	721
2016	0.50	13.90	0.20	1489
All Yrs	0.50	12.62	0.18	4190

grass percent cover likely was assimilated into other vegetation types' cover at the 250-m pixel scale and diluted as a percentage of the total at that coarser scale. The spatial difference between the 250-m pixels and AIM plots led us to conduct a  $3 \times 3$  standard deviation focal scan on our 2011 – 2016 maps to identify areas of high variability. We eliminated plots from comparison that fell in areas of high variability, with pixels of high variability defined by exceeding the median value of the focal scan. This allowed us to compare AIM plots to points on our maps where local spatial variability was moderate or low. However, the standard deviation focal scan produced results with severely reduced data ranges, which substantially affected the correlation coefficient and MAE statistics, so we rejected the standard deviation focal

scan output. In addition to the spatial resolution challenge, AIM plots can occur anywhere within a 250-m pixel. On the infrequent occasion (< 5%) when the footprint of an AIM plot fell at the intersection of multiple pixels with highly variable estimated values, we removed the plot from the analysis.

#### Leave-One-Out Validation

Leave-one-out validation is a spatially rigorous assessment technique effective at assessing maps that represent large geographical areas (Wylie et al., 2007; Zhang et al., 2011; Boyte et al., 2018). This technique mitigates spatial autocorrelation by sorting all reference data into 10 spatially and temporally randomly sorted groups and then, systematically, withholding each group as test data and using the remaining 9 groups as training data to develop 10 individual models. For this study, we chose to create 10 groups of reference data so that group members were distributed spatially, and the group test datasets were large enough to generate robust model accuracy metrics. To conduct a different analysis, we separated the 30-m annual herbaceous percent cover datasets derived from the high-resolution WorldView data from the 30-m north-central Nevada cheatgrass percent cover and the Owyhee Upland annual grass index data derived from Landsat data. We used the data derived from WorldView scenes once as test data while the data from the two other datasets were used to train the model. We then inverted the test and training datasets and repeated the process. Separating the WorldView-derived data from the north-central Nevada and the Owyhee Upland data before running the model allowed us to understand better the spatiotemporal



**Figure 3.** The 10 leave-one-out groups displayed and delineated by color. We randomly sorted the WorldView annual herbaceous vegetation (2013–2015) dataset, north-central Nevada cheatgrass percent cover dataset (~2001), and Owyhee Upland annual grass index (2006) into 10 groups and used them iteratively to evaluate model accuracy.

t4.1 **Table 4**

t4.2 The leave-one-out technique statistics. This technique sorts > 33 000 points from the  
 t4.3 WorldView, north-central Nevada, and Owyhee Uplands data into 10 random groups.  
 t4.4 We withheld each group systematically and iteratively as test data and used the remaining  
 t4.5 nine groups to train the model. In addition, we withheld all WorldView data as test data  
 t4.6 and used the Nevada and Owyhee Uplands data to train the model. We then inverted this  
 t4.7 process using the Nevada and Owyhee Uplands data as test data and the WorldView data  
 t4.8 to train the model.

t4.9	Group	Test r	Test MAE (%)	Test nRMSE (%)	Range	Test n =
t4.10	1	0.48	10.41	0.15	0-99	3960
t4.11	2	0.62	10.85	0.17	0-95	4388
t4.12	3	0.47	4.93	0.13	0-55	3221
t4.13	4	0.63	10.45	0.15	0-100	3615
t4.14	5	0.93	8.14	0.12	0-99	2208
t4.15	6	0.86	10.01	0.14	0-100	2212
t4.16	7	0.67	8.20	0.13	0-100	5659
t4.17	8	0.65	5.70	0.12	0-67	2203
t4.18	9	0.41	5.31	0.10	0-99	2639
t4.19	10	0.61	8.01	0.13	0-98	3361
t4.20	All groups	0.71	8.43	0.13	–	
t4.21	Peterson test	0.68	16.10	0.18	0-85	9018
t4.22	WorldView test	0.52	10.59	0.25	0-100	24359

421 dynamics of these datasets. For each individual model run, the  
 422 regression-tree software generated a correlation coefficient and an  
 423 MAE rate, and we calculated the nRMSE (Table 4). The leave-one-out  
 424 technique evaluated the mapped estimates of annual grass percent  
 425 cover using 250-m spatial resolution data (i.e., the WorldView-derived  
 426 data and the north-central Nevada and Owyhee Uplands datasets after  
 427 they were spatially averaged from 30 m to 250 m). The random sorting  
 428 of the dependent variable datasets generated groups that were both  
 429 spatially and temporally diverse, which helped mitigate model extrap-  
 430 olation bias.

#### 431 *Quantifying Bias*

432 We quantified bias between the annual grass maps and the leave-  
 433 one-out validation dataset. The bias was quantified at 5% and 95%  
 434 cover and determined on the basis of a theoretical minimum and max-  
 435 imum of 0 – 100% cover, which also was the range of the predicted  
 436 values. For AIM data, we used the regression equation that compared  
 437 the AIM dataset to mapped estimates of annual grass percent cover  
 438 (see Eq. (2)) at the values of 5% and 95% cover. It is important to note  
 439 that because the “any-hit” collection strategy used to gather the AIM  
 440 data can theoretically lead to data values > 100%, and the sampling strat-  
 441 egies used to create the mapped estimates’ reference data cannot gener-  
 442 ate values that exceed 100%, the calculated difference between the  
 443 estimates of annual grass percent cover and the expected y-values  
 444 does not measure true bias.

$$445 \quad y = 1.1053x + 5.4649 \quad (3)$$

446 We used the regression equation from the leave-one-out validation  
 447 data (see Eq. (1)) that was compared with the mapped estimates of an-  
 448 nual grass percent cover. This equation was used to calculate the ex-  
 449 pected y-values at the estimated values of 5% and 95% cover. The  
 450 difference between the estimates of annual grass percent cover and  
 451 the expected y-values measured bias.

$$452 \quad y = 0.8975x + 3.549 \quad (4)$$

454

#### 455 **Results**

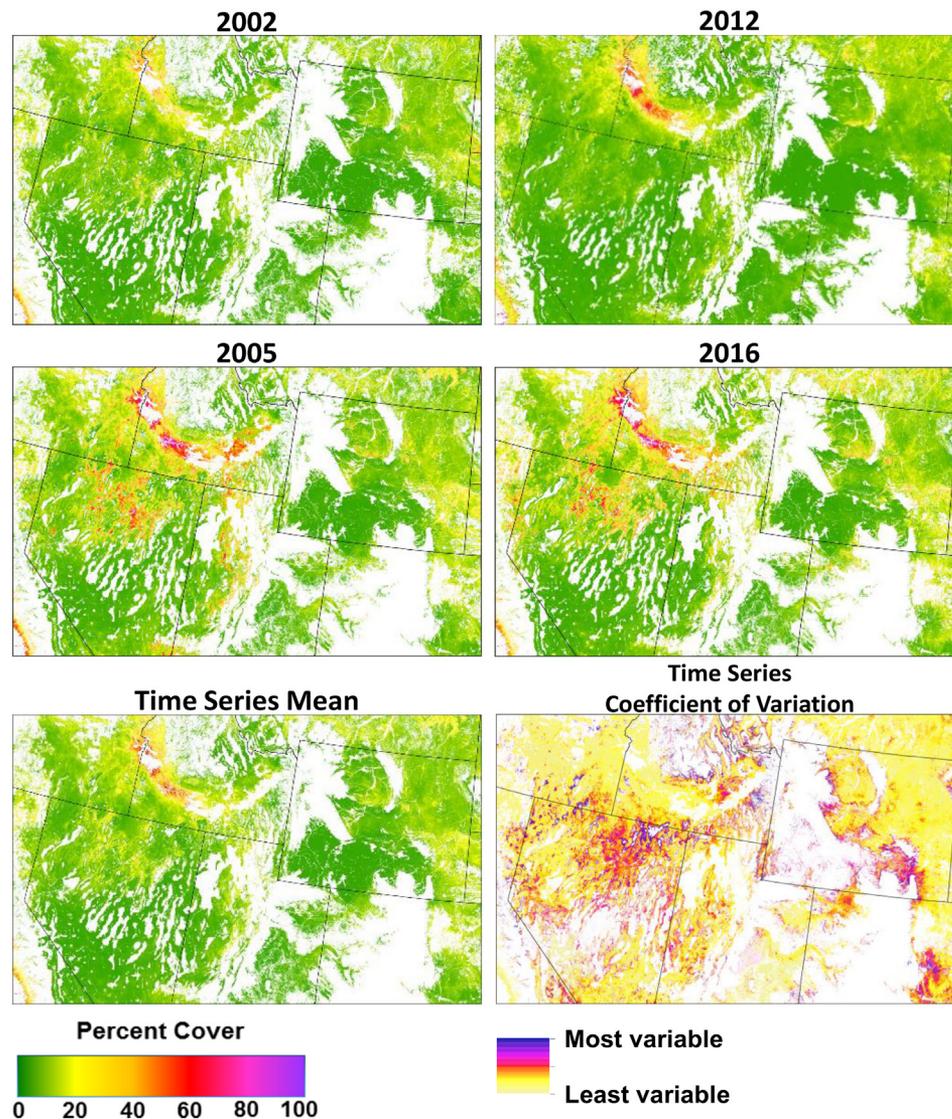
456 The model’s use of independent variables helps the user understand  
 457 the variables that most influence model development (see Table 1).  
 458 Overall, the annual grass – percent cover model uses elevation and the  
 459 Spring GSN substantially more than other variables. Summer GSN,

while seldom used to establish conditions that stratify rules, is used in 460  
 93% of the model algorithms, equal to the model’s usage of elevation 461  
 for algorithm development. Other variables the model uses frequently 462  
 include 30-yr precipitation and available water capacity, a measure of 463  
 soil water potential. Soil organic matter and SOST contribute relatively 464  
 infrequently to rules stratification, but the model uses each to develop 465  
 at least 73% of the algorithms. Topographical variables other than eleva- 466  
 tion (i.e., north steep slope and south steep slope) are used least by the 467  
 model and only for algorithm development. 468

#### 469 *Annual Grass – Percent Cover Model and Maps*

The annual grass – percent cover model shows a training and test 470  
 data  $R^2$  equal to 0.76 and 0.74, respectively (see Table 1). The model 471  
 also shows a training and test data MAE rate equal to 5.79% with a stan- 472  
 dard deviation (SD) of  $\pm 0.03$  and 5.84% with an SD of  $\pm 0.11$ , respec- 473  
 tively. The 10-fold cross validation produced an  $R^2$  of 0.74 and an MAE 474  
 rate of 5.90%. Estimated annual grass – percent cover during the 17-yr 475  
 time series ranges from 0 to 100 and experiences substantial temporal 476  
 and spatial variability. We illustrate this variability by displaying the an- 477  
 nual grass maps with the two lowest (2002 and 2012) and the two 478  
 highest (2005 and 2016) overall average percent cover in the time se- 479  
 ries (Fig. 4; Table 5); (access the time series of data and associated meta- 480  
 data at [dataset] Boyte and Wylie, 2017). The Snake River Plain 481  
 ecoregion tracks from eastern Idaho through south-central and south- 482  
 western Idaho into eastern Oregon, and this ecoregion consistently ex- 483  
 periences some of the highest annual grass percent cover in the study 484  
 area (Fig. 5). The ecoregion’s 17-yr mean annual grass percent cover 485  
 equals 19.88 compared with 7.31 for the entire study area, and 88% of 486  
 pixels that average > 60% throughout the study period are located 487  
 here. This high percent cover mapping consistency is further illustrated 488  
 by the coefficient of variation map, which shows relatively low variabil- 489  
 ity in much of this area. In northern Nevada and southeast Oregon, 2 yr – 490  
 2005 and 2016 – stand out as especially prolific for annual grass percent 491  
 cover, with values exceeding 70% in select places during both yr (see 492  
 Fig. 4). Examining the time series mean map and recognizing that the 493  
 coefficient of variation map shows relatively substantial variability in 494  
 these areas indicates that these 2 yr are likely outliers. In northern Ne- 495  
 vada during 2012, our model estimated anomalously low annual grass 496  
 percent cover following a productive 2011 (see Table 5). A relatively 497  
 substantial annual grass percent cover exists throughout much of the 498  
 study period in northeast Colorado, northwest Nebraska, and southeast 499  
 Wyoming, but overall cover in this area only occasionally exceeds 30%. 500  
 Thus, the most heavily invaded areas are northern Nevada and the 501  
 Snake River Plain, although the middle of these two areas, at the inter- 502  
 section of Nevada, Idaho, and Oregon, exhibits low annual grass percent 503  
 cover throughout the time series. 504

The annual grass percent cover datasets are each positively skewed 505  
 with substantially more low percent cover values than high percent 506  
 cover values. This is evident from the time series’ median values and 507  
 narrow 25th percentile and 75th percentile ranges while every yr’s 508  
 data range ( $\geq 99$ ) is much broader (Fig. 6A). Table 5 also demonstrates 509  
 the temporal variability of annual grass percent cover where the overall 510  
 mean percent cover ranges from 6.14 in 2012 to 11.25 in 2016. Depart- 511  
 ures from the 17-yr normal percent cover for the study period range 512  
 from 16% lower than average (2012) to 54% higher than average 513  
 (2016). Over large geographical areas, average interannual variability 514  
 in annual grass percent cover is strongly driven by weather, especially 515  
 precipitation (Bradley and Mustard, 2005; Pilliod et al., 2017), including 516  
 timing and seasonal totals (Boyte et al., 2016). Figure 6A and B illus- 517  
 trates how estimated annual grass percent cover and seasonal precipita- 518  
 tion (defined as October–May) track throughout the time series (we did 519  
 not use seasonal or annual precipitation to model annual grass percent 520  
 cover.). For most years, annual grass – percent cover patterns track 521  
 closely with seasonal precipitation. Deviations from the normal pattern 522  
 exist during 2001 and 2007 when less-than-average precipitation 523



**Figure 4.** We spatially contrast the 2 yr with the overall lowest percent cover (2002 and 2012) with the 2 yr with the overall highest percent cover (2005 and 2016). These maps are an example of the temporal variability of annual grass percent cover through the 17-yr time series. Percent cover varies annually based on disturbances, management, and weather, especially precipitation. We also show the 17-yr percent cover mean and the coefficient of variation maps. The mask (white) covers areas not classified as shrub or grassland/herbaceous by the National Land Cover Database or areas at or above 2 250-m elevation.

524 corresponds with slightly more-than-average annual grass percent cover. Also in 2008 and 2009, average and slightly less-than-average  
 525 precipitation correspond with less-than-average annual grass percent  
 526 cover. Substantial peaks in annual grass percent cover during 2005  
 527 and 2011 correspond with substantial peaks in seasonal precipitation.  
 528 The final 3 yr of the time series show progressively increasing precipita-  
 529 tion totals and progressively increasing annual grass percent cover. This  
 530 includes the highest annual grass-percent cover value of the time series  
 531 in 2016.  
 532

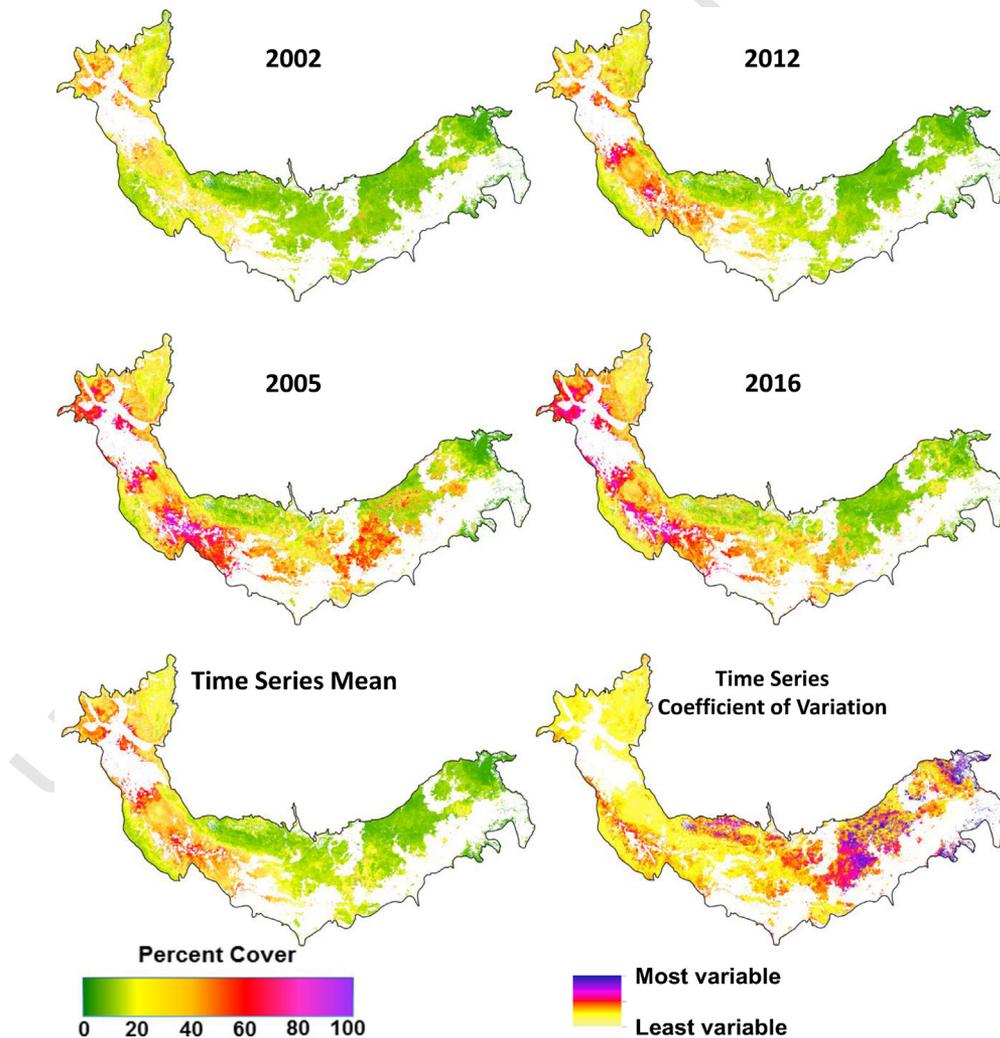
### 533 Model Evaluation

534 Despite the substantial spatial resolution differences between  
 535 datasets, a comparison of 6 yr of BLM AIM data with corresponding an-  
 536 nual grass – percent cover maps shows moderately strong agreement  
 537 for most years. With the range of yearly AIM data values being highly  
 538 variable, the correlation coefficient is marginally effective in describing  
 539 the relationship between the datasets, so we focus on the MAE rate  
 540 and nRMSE (see Table 3). We also quantified the difference at the values  
 541 of 5% and 95% cover, generally, between the AIM data and the

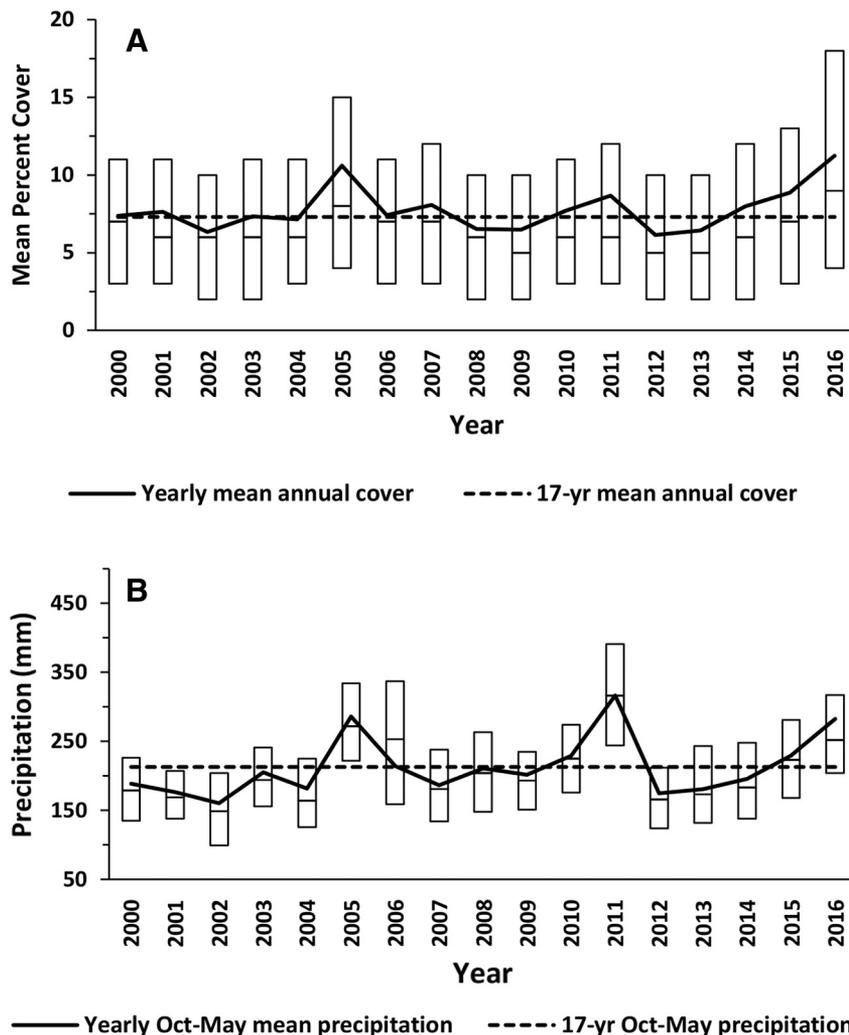
corresponding annual grass – percent cover maps for all years com- 542  
 bined. The MAE rates vary widely, with the highest rate registered in 543  
 2011 at 15.60% and the lowest rate registered in 2013 at 7.89%. The 544  
 nRMSE ranges from 0.17 in 2013 and 2014 to 0.23 in 2011. When we 545  
 combine all 6 yr of data, the MAE rate is 12.62% and the nRMSE is 546  
 0.18. Figure 7 shows the comparison between AIM data and modeled 547  
 estimates of annual grass percent cover. The scatterplot reveals an 548  
 abundance of points in the bottom left corner, a scarcity of points in 549  
 the lower right corner, and few data points above 40% on the x-axis. 550  
 Many of the more substantial differences between datasets could be re- 551  
 lated to differences in the spatial footprints of the AIM data and annual 552  
 grass – percent cover maps because variability in vegetation is likely to 553  
 be higher over the larger footprints of the maps' pixels. Overall, the data 554  
 regression line shows relative consistency with the 1:1 line. There is a 555  
 fairly consistent underestimation by the annual grass – percent cover 556  
 maps, which could be related to lower canopy hits in the AIM data. 557  
 Higher cover values than what exist in the WorldView, north-central 558  
 Nevada, and Owyhee Uplands data would result from the “any-hit” 559  
 method. At an estimated cover of 5%, the expected y-value equaled 560  
 10.99% cover based on the regression equation (Eq. (3)). At an 561

t5.1 **Table 5**  
 t5.2 Statistics describing the 17-yr time series of annual grass – percent cover estimates in the sagebrush ecosystem. The data distribution is skewed positive; therefore, the mean percent sta-  
 tistic is more sensitive than the median to high percent cover. Statistics are calculated only on unmasked areas where the National Land Cover Database classifies a pixel as shrub or grass-  
 t5.4 land/herbaceous and the elevation is at or below 2 250 m.

t5.5	Yr	Mean percent cover	Change from 17-yr mean percent cover	% change from 17-yr mean percent cover	Range
t5.6	2000	7.38	0.07	0.96	0–99
t5.7	2001	7.63	0.32	4.38	0–100
t5.8	2002	6.34	–0.97	–13.27	0–100
t5.9	2003	7.33	0.02	0.27	0–100
t5.10	2004	7.15	–0.16	–2.19	0–100
t5.11	2005	10.60	3.29	45.01	0–100
t5.12	2006	7.42	0.11	1.50	0–100
t5.13	2007	8.08	0.77	10.53	0–100
t5.14	2008	6.52	–0.79	–10.81	0–100
t5.15	2009	6.49	–0.82	–11.22	0–100
t5.16	2010	7.70	0.39	5.34	0–100
t5.17	2011	8.67	1.36	18.60	0–100
t5.18	2012	6.14	–1.17	–16.01	0–100
t5.19	2013	6.43	–0.88	–12.04	0–99
t5.20	2014	7.98	0.67	9.17	0–100
t5.21	2015	8.86	1.55	21.20	0–100
t5.22	2016	11.25	3.94	53.90	0–100
t5.23	17-yr mean	7.31	–	–	0–96



**Figure 5.** Estimated annual grass – percent cover (A) and seasonal (October–May) precipitation (B) tracked throughout the study period. The boxes show each yr’s median value and 25th and 75th percentiles. Seasonal precipitation totals were more strongly correlated to the temporal variability of estimated annual grass percent cover than were annual precipitation totals. Data were derived from data that were masked to areas classified by the National Land Cover Database as shrub or grassland/herbaceous at or below 2 250-m elevation.



**Figure 6.** Comparing all yrs of Bureau of Land Management Assessment and Inventory Monitoring data to corresponding yrs of estimated annual grass percent cover.



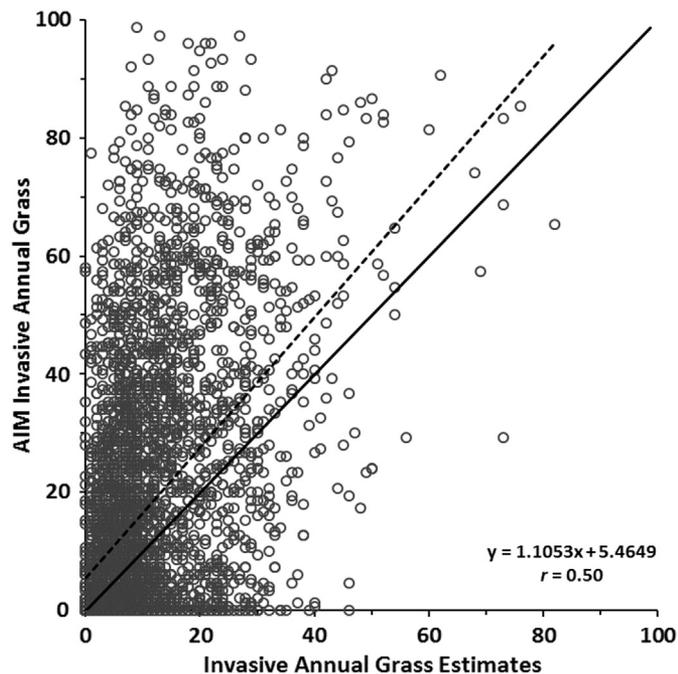
562 estimated cover of 95%, the expected  $y$ -value equaled 110.47% cover  
563 based on the regression equation.

564 The leave-one-out assessment benefits from spatial resolution  
565 agreement between datasets (i.e., both the reference data and mapped  
566 estimates of annual grass percent cover) represent a 250-m footprint.  
567 For the 10 randomly sorted groups, the lowest  $r = 0.41$  and the highest  
568  $r = 0.93$  (see Table 4). The MAE rate ranges from 4.93% to 10.85%.  
569 Combining data from the 10 groups renders an  $r$  of 0.71 and an MAE rate of  
570 8.43%. To conduct a different analysis, we separated the WorldView refer-  
571 ence data from the north-central Nevada and Owyhee Uplands refer-  
572 ence data. We used the WorldView reference data as test data,  
573 developing the model using only the north-central Nevada and Owyhee  
574 Uplands data. The model developed renders a test MAE rate of 10.59 and  
575 an nRMSE of 0.18. When the north-central Nevada and Owyhee Up-  
576 lands data are used as test data, the test MAE rate = 16.10% and  
577 nRMSE = 0.25. In the leave-one-out assessment, the location of the re-  
578 gression line in relation to the 1:1 line in Figure 8 demonstrates that the  
579 model is most accurate when it estimates about 35% annual grass per-  
580 cent cover. Below that threshold, the model slightly underestimates an-  
581 nual grass percent cover, and above that threshold the model  
582 overestimates annual grass percent cover. Quantifying bias in the  
583 leave-one-out assessment shows that an estimated cover of 5%, the ex-  
584 pected  $y$ -value equals 8.04% cover based on the regression equation  
585 (see Eq. (4)). At an estimated cover of 95%, the expected  $y$ -value equals  
586 88.81% cover.

## Discussion

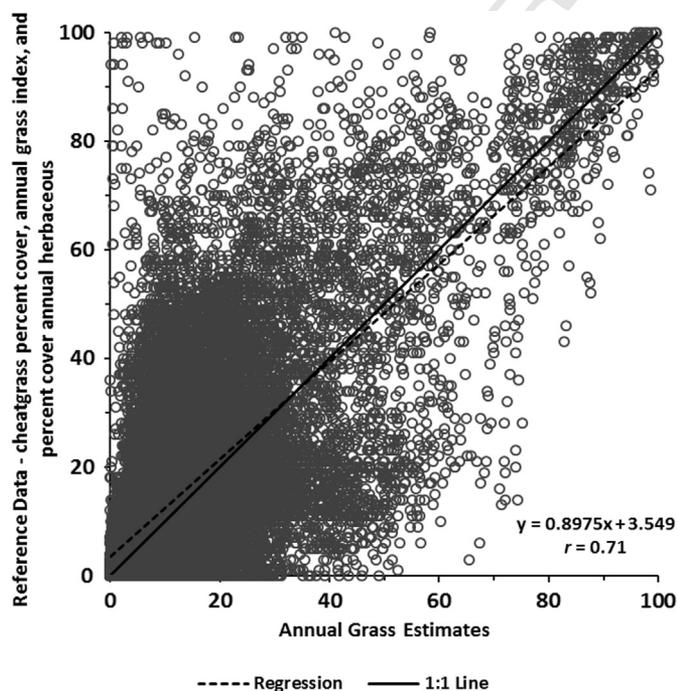
587

The invasion of annual grasses in the sagebrush ecosystem poses 588  
challenges to land managers, scientists, and practitioners because 589  
these grasses alter historical fire regimes, making fires more frequent, 590  
larger, and more severe (Whisenant, 1990; Balch et al., 2013). The 591  
changed fire regime adds to the continued degradation of the sagebrush 592  
ecosystem and requires these constituent groups to respond differently 593  
to new ecological realities. Even fighting fires has become potentially 594  
more dangerous as highly flammable grasses invade open meadows 595  
and scablands, land covers that formerly served as safe zones for fire- 596  
fighters (Kerns et al., 2016a, 2016b). The current study develops a syn- 597  
optic annual grass – percent cover model and associated maps in the 598  
sagebrush ecosystem for a 17-yr period, and this time series constitutes 599  
the longest consistent annual grass-mapping project over the broadest 600  
geographical area in the sagebrush ecosystem of which we are aware. 601  
The map results support the assertion that annual grasses are spatially 602  
variable because of site conditions (Chambers et al., 2016), even over 603  
multiple ecoregions. The map results also support the assertion that 604  
precipitation is a strong driver of annual grass percent cover (Bradley 605  
and Mustard, 2005; Pilliod et al., 2017). An increasing trend in both sea- 606  
sonal precipitation totals and annual grass percent cover is evident since 607  
2012, although the timing of precipitation could explain deviations from 608  
expected patterns of annual grass percent cover to seasonal precipita- 609  
tion totals during earlier years of the time series (see Fig. 6). 610



**Figure 7.** The leave-one-out assessment scatterplot that illustrates the relationship of AIM data to annual grass estimates of all 10 randomly sorted groups. The reference data include three spatially explicit datasets—a north-central Nevada cheatgrass percent cover dataset (~2001), an Owyhee Upland annual grass index (2006), and percent cover estimates of annual herbaceous vegetation (2013–2015).

We evaluate the model and maps using two techniques. One evaluation technique employs an approach that uses field-based BLM AIM data. We acknowledge the substantial spatial difference between the AIM data and the 250-m spatial resolution of annual grass – percent cover maps where an AIM plot covers at most 15.2% of a 250-m pixel. However, the multiyear collection strategy, ubiquity, and availability of the AIM data make it a reasonable choice for assessment of the time series, especially as continuous years of available field validation data



**Figure 8.**

that spatially match 250-m remote sensing data are extremely uncommon (Browning et al., 2015, 2017; Bradley et al., 2018). Figure 7 illustrates the relationship between data from AIM plots and corresponding estimated annual grass percent cover. Because an AIM plot is much smaller than a pixel of an annual grass – percent cover map, higher values in the AIM data will correspond with lower values in a pixel because species variability is more likely in a broader space. This makes the abundance of points in the upper left corner of the scatterplot foreseeable. The absence of all but one point in the lower right corner of Figure 7 indicates that the annual grass – percent cover model rarely produces high values at a 250-m pixel scale that are associated with low values at the AIM plot scale. This phenomenon occurring frequently would indicate significant model error, so the scarcity of such points helps validate the mapping model.

A second evaluation technique, leave-one-out, is a spatially random approach that mitigates spatial autocorrelation issues and accounts for the expansive geographic territory from where the training data are harvested. The leave-one-out comparison shows a relatively strong agreement between datasets with an overall test MAE rate of 8.43% and an overall test nRMSE of 0.13 (see Table 4). The model generates better prediction test error terms (MAE and nRMSE) when WorldView-derived data are withheld as test data than when the north-central Nevada and Owyhee Uplands data are withheld as test data. This implies that the mapping algorithms developed solely on the Great Basin and Snake River Plain areas are not as robust when they are applied to eastern extents of the study area. This may indicate functional differences between annual grass in the Great Basin and annual grass in Wyoming and Colorado. Certainly, the northeastern and southeastern parts of Wyoming will have significantly greater grass components that will complicate the separation of annual herbaceous vegetation. These east-to-west differences in annual grass – percent cover mapping can also be related to temporal differences because the north-central Nevada and Owyhee Uplands datasets are developed with reference data collected before 2007 while the WorldView datasets are developed with reference data collected after 2012.

Overall, the agreement between the dependent variable and annual grass estimates is relatively strong based on the  $r$ , MAE error rates, nRMSE, and proximity of the regression lines to the 1:1 lines in Figures 7 and 8. The difference is evident between the regression line and 1:1 line when comparing the AIM data and annual grass maps in Figure 7 and, in many cases, likely related to the spatial resolution and the “any-hit” strategy of the AIM data gathering technique. The “any-hit” effect is likely compounded at higher percent cover, which is reflected in the much higher difference quantified at 95% cover (+15.47%) than at 5% cover (+5.99%). Bias is evident in Figure 8 and demonstrates minor underestimation (–3.04%) at 5% cover, modest overestimation (+6.19%) at 95% cover, and nearly no bias at about 35% annual grass percent cover. The substantially greater number of data points at lower values than higher values in Figure 8 suggests that while annual grass can be present in high percent cover over relatively large areas (250-m pixels), in our mapped time series, low percent cover is much more likely to occur, generally reflecting variability of vegetation, litter, and bare ground components in 250-m pixels. In the leave-one-out approach, no spatial resolution difference occurs between the reference data and modeled data, so the differences between these datasets should be less than with the AIM data comparison.

We hypothesized that the time series of maps would visually exhibit both spatial and interannual variability of annual grass percent cover in already invaded areas and that this interannual variability would closely follow seasonal precipitation patterns. Areas invaded by annual grasses do experience interannual variable percent cover because of spatially dispersed and intermittent disturbances and management activities. However, weather, especially precipitation totals and timing (Bradley, 2009) and recent years’ precipitation (Pilliod et al., 2017), contributes to this interannual variability across the entire study area and, therefore, has broader impact. In Figure 6A and B, we see that 2001 and 2007 are

Q1



685 deviations from the normal mean percent cover/seasonal precipitation  
 686 pattern when less-than-average seasonal precipitation corresponds  
 687 with a slight increase in annual grass percent cover. Cheatgrass seeds  
 688 can stay viable for several years, and Pilliod et al. (2017) discovered  
 689 that, in the Great Basin, high cheatgrass seed production during 1 yr  
 690 can lead to high cheatgrass cover in subsequent years if adequate pre-  
 691 cipitation is received. We do not have data for all 3 yr before 2001, but  
 692 we do for 2007, and 2005 experienced precipitation 34% above the 17-  
 693 yr normal and annual grass percent cover 45% above the 17-yr average.  
 694 We postulate that the effects from 2005 could drive the 2007 annual  
 695 grass percent cover higher than the 17-yr average even while 2007 pre-  
 696 cipitation is slightly less than the 17-yr precipitation normal. This as-  
 697 sumes that the slightly less than normal precipitation that occurs in  
 698 2007 meets the definition of adequate (Pilliod et al., 2017). On a more  
 699 localized scale, our model estimates anomalously low annual grass per-  
 700 cent cover in northern Nevada in 2012 following a highly productive  
 701 2011 for this area. This phenomenon is not necessarily unexpected,  
 702 even though a prolific growing season would leave abundant seed for  
 703 germination during the subsequent growing season (Kerns and Day,  
 704 2017). This is because adequate moisture must be present at the correct  
 705 time before seeds can germinate. A progressive increase in annual grass  
 706 percent cover can be observed in northern Nevada from 2013 until 2016  
 707 (see Fig. 4). This localized area might serve as a harbinger of high overall  
 708 percent cover for the entire study area. In 3 of the 4 highest yr of overall  
 709 annual grass percent cover (see Table 5), this localized area also has its  
 710 highest annual grass percent cover. In no other year does this geographi-  
 711 cal area show the level of invasion it does during 2005, 2011, and 2016  
 712 (see Fig. 4; Table 5).

713 Most of the study area shows annual grass percent cover during  
 714 some, if not all, years in our time series. However, some areas are typi-  
 715 cally void of substantial annual grass production (e.g., the southwest  
 716 corner of Wyoming, much of central Nevada, and east-central Califor-  
 717 nia). It is likely that these areas possess characteristics that allow them  
 718 to resist annual grass dominance. Chambers et al. (2016) discusses sev-  
 719 eral environmental characteristics that drive plant community resis-  
 720 tance to exotic annual *Brome* grass invasion including elevation,  
 721 climate, and soils. Our annual grass model's usage of variables converges  
 722 with these findings where elevation represents the second most heavily  
 723 used driver (see Table 1), and soil metrics—available water capacity and  
 724 soil organic matter—and 30-yr precipitation are used heavily as well.  
 725 Other characteristics can contribute to the resistance of annual grass in-  
 726 vasion including the presence of intact biotic soil crusts (Condon and  
 727 Pyke, 2018), traits of invading plants, interactions between the invaders,  
 728 native plant communities (Chambers et al., 2014, 2016), and land use  
 729 history (Pyke et al., 2016).

## 730 Implications

731 Annual grasses in the sagebrush ecosystem present challenges to the  
 732 responsible use and management of this ecosystem. The more substan-  
 733 tial the annual grass domination, the more fire regimes change and in-  
 734 crease the severity and frequency of disturbances, and the more  
 735 difficult it becomes to manage for multiple uses like wildlife habitat, re-  
 736 creation, grazing, and development. Management needs to be able to  
 737 identify actions that may alleviate the current cycle of invasive annual  
 738 grass species. The time series of maps developed for this study allows  
 739 examination of annual grass distribution trends and, through these  
 740 trends, supports the understanding that annual grass percent cover ex-  
 741 periences spatiotemporal variability in this ecosystem for specific rea-  
 742 sons. Understanding the drivers of these dynamics provides land  
 743 managers, scientists, and practitioners with the tools needed to better  
 744 understand, manage, and use the ecosystem, especially if they under-  
 745 stand when and to what degree these drivers are likely to change. The  
 746 time series allows longitudinal comparisons with temporally dynamic  
 747 exogenous conditions like weather and grazing. Therefore, correlating  
 748 annual grass percent cover variability throughout the time series with

changing exogenous conditions could help predict future conditions. 749  
 The time series also allows comparisons with spatially endogenous con- 750  
 ditions like topography, soil characteristics, and species competition. 751

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